

Modeling Moderated Variance Ratios in Mplus

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Introduction

The Mplus package (Muthén and Muthén, 2011) allows restrictions to be imposed on variables that are functions of variances and covariances. Thus, for a categorical moderator, the equal variances ratio (EVR) test described in Smithson (2012) can be incorporated into a multi-groups structural equations model (SEM) via appropriate model comparisons. These SEMs also can incorporate tests of homogeneity of (error) variance and, of course, moderation of correlations and moderation of slopes.

EVR models are implemented in correlation SEMs of the kind described in Smithson (2012). The EVR component is initiated by defining the relevant variance ratios. Suppose the variance ratios of interest are for the i^{th} and j^{th} categories of the moderator. These variance ratios must first be defined in the MODEL CONSTRAINTS statement. We can then compare a model that leaves them unconstrained against a model that imposes equality between them (via a statement such as `ratioi = ratioj`). Clearly multiple hypotheses of this kind can be tested simultaneously. Model comparisons can be made in the usual fashion.

Homogeneity of variance also can be tested in the correlation SEMs, either for x or y alone (which is incompatible with EVR) or simultaneously (which imposes EVR). Regression SEMs can include tests for homogeneity of error variance, and regression SEMs also can incorporate EVR tests. All of these capabilities are demonstrated in the examples that follow.

Examples of EVR-testing SEMs in Mplus

Two-Category Moderator

This is a two-category moderator example in which the null hypothesis of EVR is true. It uses an artificial data-set sampled from two bivariate normal distributions, associated with one category of a binary moderator variable z which takes values -1 and +1. For the first moderator category $\sigma_x^2 = 1$ and $\sigma_y^2 = 2$, while for the second category $\sigma_x^2 = 4$ and $\sigma_y^2 = 8$. Thus, the population variance

ratio in both moderator categories is 1/2. The population covariances are 1 and the means are 0 for both categories. The data-file is eqdat.dat. There are 500 observations in each category. In the first category, $s_x^2 = 0.883$ and $s_y^2 = 1.958$, so $s_x^2/s_y^2 = 0.451$. In the second category, $s_x^2 = 4.068$ and $s_y^2 = 7.515$, so $s_x^2/s_y^2 = 0.541$.

We first fit a saturated model permitting unequal variance ratios:

```

TITLE:          EVR not imposed
DATA:           FILE = eqdat.dat;
VARIABLE:       NAMES = x y z;
                USEVARIABLES = x y;
                GROUPING IS z (-1 = cat1 1 = cat2);
MODEL:          fx by x* (sd1);
                fy by y* (sd2);
                fx@1;
                fy@1;
                x@0;
                y@0;
                fx with fy;
MODEL cat2:
                fx by x* (sd3);
                fy by y* (sd4);
MODEL CONSTRAINT:
                NEW(ratio1, ratio2);
                ratio1 = sd1^2/sd2^2;
                ratio2 = sd3^2/sd4^2;

```

The output reproduces the sample variance ratios, 0.451 and 0.541.

```

SUMMARY OF ANALYSIS
Number of groups                2
Number of observations
  Group CAT1                    500
  Group CAT2                    500
Number of dependent variables   2
Number of independent variables 0
Number of continuous latent variables 2
Observed dependent variables
  Continuous
    X          Y
Continuous latent variables
  FX          FY
Variables with special functions
  Grouping variable    Z
Estimator                                ML
Information matrix          OBSERVED
Maximum number of iterations 1000

```

Convergence criterion 0.500D-04
 Maximum number of steepest descent iterations 20
 Input data file(s)
 eqdat.dat
 Input data format FREE
 THE MODEL ESTIMATION TERMINATED NORMALLY

MODEL FIT INFORMATION

Number of Free Parameters 10
 Loglikelihood
 H0 Value -3626.178
 H1 Value -3626.178
 Information Criteria
 Akaike (AIC) 7272.356
 Bayesian (BIC) 7321.433
 Sample-Size Adjusted BIC 7289.673
 (n* = (n + 2) / 24)
 Chi-Square Test of Model Fit
 Value 0.000
 Degrees of Freedom 0
 P-Value 0.0000
 Chi-Square Contributions From Each Group
 CAT1 0.000
 CAT2 0.000
 RMSEA (Root Mean Square Error Of Approximation)
 Estimate 0.000
 90 Percent C.I. 0.000 0.000
 Probability RMSEA <= .05 0.000
 CFI/TLI
 CFI 1.000
 TLI 1.000
 Chi-Square Test of Model Fit for the Baseline Model
 Value 403.135
 Degrees of Freedom 2
 P-Value 0.0000
 SRMR (Standardized Root Mean Square Residual)
 Value 0.000

MODEL RESULTS

Group		Estimate	S.E.	Est./S.E.	Two-Tailed P-Value
CAT1					
FX	BY				
	X	0.939	0.030	31.623	0.000
FY	BY				
	Y	1.398	0.044	31.623	0.000
FX	WITH				

	FY	0.737	0.020	36.054	0.000	
Means						
	FX	0.000	0.000	999.000	999.000	
	FY	0.000	0.000	999.000	999.000	
Intercepts						
	X	0.033	0.042	0.774	0.439	
	Y	0.049	0.063	0.784	0.433	
Variances						
	FX	1.000	0.000	999.000	999.000	
	FY	1.000	0.000	999.000	999.000	
Residual Variances						
	X	0.000	0.000	999.000	999.000	
	Y	0.000	0.000	999.000	999.000	
Group CAT2						
	FX	BY				
	X		2.015	0.064	31.623	0.000
	FY	BY				
	Y		2.739	0.087	31.623	0.000
	FX	WITH				
	FY		0.152	0.044	3.472	0.001
Means						
	FX		-0.020	0.049	-0.410	0.682
	FY		-0.037	0.050	-0.732	0.464
Intercepts						
	X		0.033	0.042	0.774	0.439
	Y		0.049	0.063	0.784	0.433
Variances						
	FX		1.000	0.000	999.000	999.000
	FY		1.000	0.000	999.000	999.000
Residual Variances						
	X		0.000	0.000	999.000	999.000
	Y		0.000	0.000	999.000	999.000
New/Additional Parameters						
	RATIO1		0.451	0.027	16.538	0.000
	RATIO2		0.541	0.048	11.311	0.000

The next model restricts the variance ratios to be equal:

```

TITLE:          This is EVR, equal variance ratios
DATA:          FILE = eqdat.dat;
VARIABLE:      NAMES = x y z;
               USEVARIABLES = x y;
               GROUPING IS z (-1 = cat1 1 = cat2);
MODEL:        fx by x* (sdx1);
               fy by y* (sdy1);
               fx@1;

```

```

        fy@1;
        x@0;
        y@0;
        fx with fy;
MODEL cat2:
    fx by x* (sdx2);
    fy by y* (sdy2);
MODEL CONSTRAINT:
    NEW(ratio1, ratio2);
    ratio1 = sdx1^2/sdy1^2;
    ratio2 = sdx2^2/sdy2^2;
    ratio1 = ratio2;

```

Its output shows that this model cannot be rejected ($\chi^2(1) = 2.920, p = .0875$), and estimates the variance-ratio to be 0.478.

```

MODEL FIT INFORMATION
Number of Free Parameters          9
Loglikelihood
    H0 Value                       -3627.638
    H1 Value                       -3626.178
Information Criteria
    Akaike (AIC)                   7273.276
    Bayesian (BIC)                 7317.446
    Sample-Size Adjusted BIC       7288.861
    (n* = (n + 2) / 24)
Chi-Square Test of Model Fit
    Value                           2.920
    Degrees of Freedom              1
    P-Value                         0.0875
Chi-Square Contributions From Each Group
    CAT1                            0.931
    CAT2                            1.989
RMSEA (Root Mean Square Error Of Approximation)
    Estimate                        0.062
    90 Percent C.I.                0.000  0.150
    Probability RMSEA <= .05       0.280
CFI/TLI
    CFI                            0.995
    TLI                            0.990
Chi-Square Test of Model Fit for the Baseline Model
    Value                           403.135
    Degrees of Freedom              2
    P-Value                         0.0000
SRMR (Standardized Root Mean Square Residual)
    Value                           0.031
MODEL RESULTS

```

	Estimate	S.E.	Est./S.E.	Two-Tailed P-Value
Group CAT1				
FX				
BY				
X	0.953	0.029	32.672	0.000
FY				
BY				
Y	1.378	0.042	32.995	0.000
FX				
WITH				
FY	0.737	0.020	35.987	0.000
Means				
FX	0.000	0.000	999.000	999.000
FY	0.000	0.000	999.000	999.000
Intercepts				
X	0.033	0.043	0.767	0.443
Y	0.049	0.062	0.800	0.424
Variances				
FX	1.000	0.000	999.000	999.000
FY	1.000	0.000	999.000	999.000
Residual Variances				
X	0.000	0.000	999.000	999.000
Y	0.000	0.000	999.000	999.000
Group CAT2				
FX				
BY				
X	1.955	0.050	39.256	0.000
FY				
BY				
Y	2.828	0.074	38.129	0.000
FX				
WITH				
FY	0.151	0.044	3.465	0.001
Means				
FX	-0.021	0.050	-0.421	0.673
FY	-0.036	0.050	-0.719	0.472
Intercepts				
X	0.033	0.043	0.767	0.443
Y	0.049	0.062	0.800	0.424
Variances				
FX	1.000	0.000	999.000	999.000
FY	1.000	0.000	999.000	999.000
Residual Variances				
X	0.000	0.000	999.000	999.000
Y	0.000	0.000	999.000	999.000
New/Additional Parameters				
RATIO1	0.478	0.024	19.996	0.000
RATIO2	0.478	0.024	19.996	0.000

Regression SEMs also can incorporate EVR tests. The variance ratios have

to be derived from the error-variances, whence the definitions of `ratio1` and `ratio2` in the syntax below.

```

TITLE:          This is EVR, equal variance ratios
DATA:           FILE = eqdat.dat;
VARIABLE:       NAMES = x y z;
                USEVARIABLES = x y;
                GROUPING IS z (-1 = cat1 1 = cat2);
MODEL:          x* (sdx1);
                y* (se1);
                y on x (b1);
MODEL cat2:
                x* (sdx2);
                y* (se2);
                y on x (b2);
MODEL CONSTRAINT:
                NEW(ratio1, ratio2);
                ratio1 = sdx1/(se1 + (sdx1)*(b12));
                ratio2 = sdx2/(se2 + (sdx2)*(b22));
                ratio1 = ratio2;

```

This type of model provides estimates of variances rather than standard deviations, which is slightly confusing when comparing its estimates with those of the correlation model. Nonetheless, its estimates are consistent with those of the correlation model. Indeed, the ratio of the estimated correlations in the correlation model and regression coefficients in this model are identical (up to roundoff error): $.737/.151 = 0.488$ and $1.066/.219 = 0.487$.

MODEL RESULTS

		Estimate	S.E.	Est./S.E.	Two-Tailed P-Value
Group CAT1					
Y	ON				
	X	1.066	0.040	26.408	0.000
Means					
	X	0.033	0.043	0.767	0.443
Intercepts					
	Y	0.014	0.042	0.347	0.729
Variances					
	X	0.907	0.056	16.336	0.000
Residual Variances					
	Y	0.869	0.052	16.682	0.000
Group CAT2					
Y	ON				
	X	0.219	0.063	3.454	0.001
Means					

X	-0.008	0.087	-0.095	0.924
Intercepts				
Y	-0.050	0.125	-0.401	0.689
Variances				
X	3.822	0.195	19.628	0.000
Residual Variances				
Y	7.815	0.410	19.051	0.000
New/Additional Parameters				
RATIO1	0.478	0.024	19.996	0.000
RATIO2	0.478	0.024	19.996	0.000

SEM Example in Smithson (2012)

The data-set for this example is as described in Smithson (2012), and the file is semex.dat. Starting with SEMs for regression coefficients, we begin with a model that assumes HoV for X but otherwise leaves the other parameters free.

```

TITLE:          HoV in X SEM example
DATA:          FILE = semex.dat;
VARIABLE:      NAMES = x y z;
               USEVARIABLES = x y;
               GROUPING IS z (-1 = cat1 1 = cat2);
MODEL:         x* (sdx1);
               y* (se1);
               y on x (b1);
MODEL cat2:
               x* (sdx2);
               y* (se2);
               y on x (b2);
MODEL CONSTRAINT:
               NEW(ratio1, ratio2);
               ratio1 = sdx1/(se1 + (sdx1)*(b12));
               ratio2 = sdx2/(se2 + (sdx2)*(b22));
               sdx1 = sdx2;

```

As reported in Smithson (2012), the model fit is $\chi^2(1) = 0.370$ ($p = .543$). Next, we estimate a model that imposes equal variances ratios.

```

DATA:          FILE = semex.dat;
VARIABLE:      NAMES = x y z;
               USEVARIABLES = x y;
               GROUPING IS z (-1 = cat1 1 = cat2);
MODEL:         x* (sdx1);
               y* (se1);
               y on x (b1);
MODEL cat2:
               x* (sdx2);

```



```

        y* (se2);
        y on x (b2);
MODEL CONSTRAINT:
        NEW(ratio1, ratio2);
        ratio1 = sdx1/(se1 + (sdx1)*(b1^2));
        ratio2 = sdx2/(se2 + (sdx2)*(b2^2));
        ratio1 = ratio2;

```

This model yields a chi-square fit of $\chi^2(1) = 82.246$ ($p < .0001$), so we may reject the EVR hypothesis. We then fit a regression model that relaxes EVR but forces equality of regression coefficients and HoV in X .

```

TITLE:          HoV in X and equal slopes SEM example
DATA:          FILE = semex.dat;
VARIABLE:      NAMES = x y z;
                USEVARIABLES = x y;
                GROUPING IS z (-1 = cat1 1 = cat2);
MODEL:         x* (sdx1);
                y* (se1);
                y on x (b1);
MODEL cat2:
                x* (sdx2);
                y* (se2);
                y on x (b2);
MODEL CONSTRAINT:
                sdx1 = sdx2;
                b1 = b2;

```

This model can be rejected because $\chi^2(2) = 15.779$ ($p = .0004$), and the fit comparison between this model and the HoV in X model yields $\chi^2(1) = 15.779 - 0.370 = 15.429$ $p < .0001$, suggesting significantly worse fit. So we conclude that there is moderation of slopes.

Turning now to SEMs for correlations, as before we begin with a model that assumes HoV for X but otherwise leaves the other parameters free.

```

TITLE:          HoV in X SEM example
DATA:          FILE = semex.dat;
VARIABLE:      NAMES = x y z;
                USEVARIABLES = x y;
                GROUPING IS z (-1 = cat1 1 = cat2);
MODEL:         fx by x* (sdx1);
                fy by y* (sdy1);
                fx@1;
                fy@1;
                x@0;
                y@0;
                fx with fy;

```

```

MODEL cat2:
    fy by y* (sdy2);

```

As reported in Smithson (2012), the model fit is $\chi^2(1) = 0.370$ ($p = .543$). Next, a model restricting the variance ratios to equality is estimated:

```

TITLE:          EVR SEM example
DATA:           FILE = semex.dat;
VARIABLE:       NAMES = x y z;
                USEVARIABLES = x y;
                GROUPING IS z (-1 = cat1 1 = cat2);
MODEL:          fx by x* (sd1);
                fy by y* (sdy1);
                fx@1;
                fy@1;
                x@0;
                y@0;
                fx with fy;
MODEL cat2:
    fx by x* (sd2);
    fy by y* (sdy2);
MODEL CONSTRAINT:
    NEW(ratio1, ratio2);
    ratio1 = sd1^2/sdy1^2;
    ratio2 = sd2^2/sdy2^2;
    ratio1 = ratio2;

```

As with its regression counterpart, this model yields a chi-square fit of $\chi^2(1) = 82.246$ ($p < .0001$). Finally, we estimate an equal-correlation model.

```

TITLE:          HoV in X and equal correlations SEM example
DATA:           FILE = semex.dat;
VARIABLE:       NAMES = x y z;
                USEVARIABLES = x y;
                GROUPING IS z (-1 = cat1 1 = cat2);
MODEL:          fx by x* (sd1);
                fy by y* (sdy1);
                fx@1;
                fy@1;
                x@0;
                y@0;
                fx with fy (1);
MODEL cat2:
    fy by y* (sdy2);

```

We have $\chi^2(2) = 0.453$ ($p = .797$). The model comparison result is $\chi^2(1) = 0.083$ ($p = .773$), suggesting that the equal-correlations model is a good fit to the data.

Four-Category Moderator

Now we consider a four-category moderator example, with EVR for the first two categories and for the second two, but not for both pairs of categories. We generate an artificial data-set sampled from four bivariate normal distributions, associated with a moderator variable z which takes values 1, 2, 3, and 4, with 300 observations in each category. The variance ratio for the first two moderator categories equals 1/2, whereas for the third and fourth categories the ratio is 1/6. The first two categories' correlations also are identical. The data-set is named fourdat.dat. We begin with a model that tests both EVR hypotheses and the equal-correlations hypothesis.

```
TITLE:          EVR and equal correlations hypotheses: four groups example
DATA:          FILE = fourdat.dat;
VARIABLE:      NAMES = x y group;
               USEVARIABLES = x y;
               GROUPING IS group (1 = cat1 2 = cat2 3 = cat3 4 = cat4);
MODEL:        fx by x* (sdx1);
               fy by y* (sdy1);
               fx@1;
               fy@1;
               x@0;
               y@0;
               fx with fy* (r1);
MODEL cat2:
               fx by x* (sdx2);
               fy by y* (sdy2);
               fx with fy* (r2);
MODEL cat3:
               fx by x* (sdx3);
               fy by y* (sdy3);
               fx with fy* (r3);
MODEL cat4:
               fx by x* (sdx4);
               fy by y* (sdy4);
               fx with fy* (r4);
MODEL CONSTRAINT:
               NEW(ratio1, ratio2, ratio3, ratio4);
               ratio1 = sdx12/sdy12;
               ratio2 = sdx22/sdy22;
               ratio3 = sdx32/sdy32;
               ratio4 = sdx42/sdy42;
               ratio1 = ratio2;
               ratio3 = ratio4;
               r1 = r2;
```

The results are shown below. The model cannot be rejected ($\chi^2(3) = 2.250, p = .522$), and estimates the variance-ratios to be 0.506 and 0.176.

SUMMARY OF ANALYSIS

Number of groups		4
Number of observations		
Group CAT1		300
Group CAT2		300
Group CAT3		300
Group CAT4		300
Number of dependent variables		2
Number of independent variables		0
Number of continuous latent variables		2
Observed dependent variables		
Continuous		
X	Y	
Continuous latent variables		
FX	FY	
Variables with special functions		
Grouping variable	GROUP	
Estimator		ML
Information matrix		OBSERVED
Maximum number of iterations		1000
Convergence criterion		0.500D-04
Maximum number of steepest descent iterations		20
Input data file(s)		
fourdat.dat		
Input data format	FREE	
THE MODEL ESTIMATION TERMINATED NORMALLY		
MODEL FIT INFORMATION		
Number of Free Parameters		17
Loglikelihood		
H0 Value		-4021.155
H1 Value		-4020.031
Information Criteria		
Akaike (AIC)		8076.311
Bayesian (BIC)		8162.842
Sample-Size Adjusted BIC		8108.844
(n* = (n + 2) / 24)		
Chi-Square Test of Model Fit		
Value		2.250
Degrees of Freedom		3
P-Value		0.5222
Chi-Square Contributions From Each Group		
CAT1		1.122
CAT2		1.122

CAT3	0.001	
CAT4	0.004	
RMSEA (Root Mean Square Error Of Approximation)		
Estimate	0.000	
90 Percent C.I.	0.000	0.087
Probability RMSEA <= .05	0.764	
CFI/TLI		
CFI	1.000	
TLI	1.001	
Chi-Square Test of Model Fit for the Baseline Model		
Value	904.400	
Degrees of Freedom	4	
P-Value	0.0000	
SRMR (Standardized Root Mean Square Residual)		
Value	0.022	

MODEL RESULTS

	Estimate	S.E.	Est./S.E.	Two-Tailed P-Value
Group CAT1				
FX BY				
X	1.025	0.036	28.517	0.000
FY BY				
Y	1.441	0.050	28.803	0.000
FX WITH				
FY	0.734	0.019	38.935	0.000
Means				
FX	0.000	0.000	999.000	999.000
FY	0.000	0.000	999.000	999.000
Intercepts				
X	-0.040	0.059	-0.671	0.502
Y	-0.082	0.083	-0.983	0.325
Variances				
FX	1.000	0.000	999.000	999.000
FY	1.000	0.000	999.000	999.000
Residual Variances				
X	0.000	0.000	999.000	999.000
Y	0.000	0.000	999.000	999.000

Group CAT2				
FX BY				
X	1.231	0.044	28.015	0.000
FY BY				
Y	1.731	0.062	27.729	0.000
FX WITH				
FY	0.734	0.019	38.935	0.000
Means				

FX		0.031	0.075	0.419	0.675
FY		0.064	0.075	0.854	0.393
Intercepts					
X		-0.040	0.059	-0.671	0.502
Y		-0.082	0.083	-0.983	0.325
Variances					
FX		1.000	0.000	999.000	999.000
FY		1.000	0.000	999.000	999.000
Residual Variances					
X		0.000	0.000	999.000	999.000
Y		0.000	0.000	999.000	999.000
Group CAT3					
FX	BY				
X		1.005	0.040	25.061	0.000
FY	BY				
Y		2.395	0.096	25.054	0.000
FX	WITH				
FY		0.833	0.018	46.979	0.000
Means					
FX		0.045	0.082	0.545	0.586
FY		0.090	0.067	1.339	0.180
Intercepts					
X		-0.040	0.059	-0.671	0.502
Y		-0.082	0.083	-0.983	0.325
Variances					
FX		1.000	0.000	999.000	999.000
FY		1.000	0.000	999.000	999.000
Residual Variances					
X		0.000	0.000	999.000	999.000
Y		0.000	0.000	999.000	999.000
Group CAT4					
FX	BY				
X		1.393	0.049	28.622	0.000
FY	BY				
Y		3.321	0.116	28.648	0.000
FX	WITH				
FY		0.495	0.044	11.372	0.000
Means					
FX		0.051	0.072	0.714	0.475
FY		0.110	0.063	1.738	0.082=
Intercepts					
X		-0.040	0.059	-0.671	0.502
Y		-0.082	0.083	-0.983	0.325
Variances					

FX	1.000	0.000	999.000	999.000
FY	1.000	0.000	999.000	999.000
Residual Variances				
X	0.000	0.000	999.000	999.000
Y	0.000	0.000	999.000	999.000
New/Additional Parameters				
RATIO1	0.506	0.028	18.032	0.000
RATIO2	0.506	0.028	18.032	0.000
RATIO3	0.176	0.009	18.541	0.000
RATIO4	0.176	0.009	18.541	0.000

We then estimate a model that constrains all variance ratios to be equal. The syntax is shown below. This model can be rejected ($\chi^2(3) = 176.254, p < .0001$).

```

TITLE:          EVR four groups example
DATA:          FILE = fourdat.dat;
VARIABLE:      NAMES = x y group;
               USEVARIABLES = x y;
               GROUPING IS group (1 = cat1 2 = cat2 3 = cat3 4 = cat4);
MODEL:         fx by x* (sdx1);
               fy by y* (sdy1);
               fx@1;
               fy@1;
               x@0;
               y@0;
               fx with fy* (r1);
MODEL cat2:
               fx by x* (sdx2);
               fy by y* (sdy2);
               fx with fy* (r2);
MODEL cat3:
               fx by x* (sdx3);
               fy by y* (sdy3);
               fx with fy* (r3);
MODEL cat4:
               fx by x* (sdx4);
               fy by y* (sdy4);
               fx with fy* (r4);
MODEL CONSTRAINT:
               NEW(ratio1, ratio2, ratio3, ratio4);
               ratio1 = sdx12/sdy12;
               ratio2 = sdx22/sdy22;
               ratio3 = sdx32/sdy32;
               ratio4 = sdx42/sdy42;
               ratio1 = ratio2;
               ratio1 = ratio3;
               ratio1 = ratio4;

```

As noted earlier, is also possible to incorporate EVR tests in a regression SEM. In this example, the hypothesis that the correlations in the first two categories are identical is equivalent to the hypothesis that the corresponding regression coefficients are identical. The following syntax extracts the EVR hypotheses from the error-variances and variances for X .

```

TITLE:          EVR and equal slopes hypotheses: four groups example
DATA:          FILE = fourdat.dat;
VARIABLE:      NAMES = x y group;
               USEVARIABLES = x y;
               GROUPING IS group (1 = cat1 2 = cat2 3 = cat3 4 = cat4);
MODEL:        x* (sdx1);
               y* (se1);
               y on x* (b1);
MODEL cat2:   x* (sdx2);
               y* (se2);
               y on x* (b2);
MODEL cat3:   x* (sdx3);
               y* (se3);
               y on x* (b3);
MODEL cat4:   x* (sdx4);
               y* (se4);
               y on x* (b4);
MODEL CONSTRAINT:
               NEW(ratio1, ratio2, ratio3, ratio4);
               ratio1 = sdx1/(se1 + sdx1*b12);
               ratio2 = sdx2/(se2 + sdx2*b22);
               ratio3 = sdx3/(se3 + sdx3*b32);
               ratio4 = sdx4/(se4 + sdx4*b42);
               ratio1 = ratio2;
               ratio3 = ratio4;
               b1 = b2;

```

The output is shown next. As noted earlier this type of model provides estimates of variances rather than standard deviations, but the estimates are consistent with those in the correlation model. Most importantly, due to the EVR for the third and fourth categories of the moderator, we should expect the ratios of the estimated correlations and regression coefficients in those two groups to be the same, and indeed they are (within roundoff error): $.833/.495 = 1.984/1.181 = 1.68$.

MODEL FIT INFORMATION	
Number of Free Parameters	17
Loglikelihood	

H0 Value	-4021.155	
H1 Value	-4020.031	
Information Criteria		
Akaike (AIC)	8076.311	
Bayesian (BIC)	8162.842	
Sample-Size Adjusted BIC	8108.844	
(n* = (n + 2) / 24)		
Chi-Square Test of Model Fit		
Value	2.250	
Degrees of Freedom	3	
P-Value	0.5222	
Chi-Square Contributions From Each Group		
CAT1	1.122	
CAT2	1.122	
CAT3	0.002	
CAT4	0.004	
RMSEA (Root Mean Square Error Of Approximation)		
Estimate	0.000	
90 Percent C.I.	0.000	0.087
Probability RMSEA <= .05	0.764	
CFI/TLI		
CFI	1.000	
TLI	1.001	
Chi-Square Test of Model Fit for the Baseline Model		
Value	904.400	
Degrees of Freedom	4	
P-Value	0.0000	
SRMR (Standardized Root Mean Square Residual)		
Value	0.022	

MODEL RESULTS

	Estimate	S.E.	Est./S.E.	Two-Tailed P-Value
Group CAT1				
Y	ON			
X	1.032	0.039	26.441	0.000
Means				
X	-0.040	0.059	-0.670	0.503
Intercepts				
Y	-0.041	0.057	-0.722	0.470
Variances				
X	1.051	0.074	14.259	0.000
Residual Variances				
Y	0.958	0.068	14.007	0.000

Group CAT2
Y ON

X	1.032	0.039	26.441	0.000
Means				
X	-0.001	0.071	-0.014	0.989
Intercepts				
Y	0.030	0.068	0.446	0.655
Variances				
X	1.516	0.108	14.008	0.000
Residual Variances				
Y	1.382	0.097	14.259	0.000
Group CAT3				
Y	ON			
X	1.984	0.068	29.129	0.000
Means				
X	0.005	0.058	0.093	0.926
Intercepts				
Y	0.124	0.077	1.617	0.106
Variances				
X	1.010	0.081	12.531	0.000
Residual Variances				
Y	1.761	0.141	12.518	0.000
Group CAT4				
Y	ON			
X	1.181	0.109	10.869	0.000
Means				
X	0.032	0.080	0.393	0.695
Intercepts				
Y	0.245	0.167	1.469	0.142
Variances				
X	1.942	0.136	14.310	0.000
Residual Variances				
Y	8.321	0.581	14.327	0.000
New/Additional Parameters				
RATIO1	0.506	0.028	18.032	0.000
RATIO2	0.506	0.028	18.032	0.000
RATIO3	0.176	0.009	18.540	0.000
RATIO4	0.176	0.009	18.540	0.000

References

Muthén, L., & Muthén, B. (2010). Mplus users guide (6th ed.). Los Angeles, CA: Muthén and Muthén.

Smithson, M. (2012). A simple statistic for comparing moderation of slopes and

correlations. The Australian National University, Canberra, Australia: Unpublished manuscript.